Rent Sharing and Inclusive Growth

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CENTRE for ECONOMIC PERFORMANCE Motivation



FIGURE 2.3: ANNUAL MEDIAN REAL WEEKLY EARNINGS

Notes: Weekly earnings deflated by CPI, CPIH and RPIJ. Source: Annual Survey of Hours and Earnings (ASHE).

FIGURE 2.2: PRODUCTIVITY, WAGES, AND COMPENSATION



Notes: Growth rates of real productivity, real compensation (deflated by the GVA deflator), real average and median wages per hour (deflated by the CPI). Source: OECD National Accounts, ONS.

LSE Growth Commission (2017)

Motivation

Harold Meyerson, American Prospect (2014)

for the vast majority of American workers, the link between their productivity and their compensation no longer exists

The Economist (2013)

unless you are rich, [gross domestic product] growth isn't doing much to raise your income anymore

Anna Stansbury and Lawrence Summers, FT (2017)

productivity growth is doing much more to raise typical pay than an initial look at the productivity-pay divergence [suggests]

- Is it about productivity of median workers or rather their bargaining power?
- ► The role of firm's wage setting process has been overlooked.

- The long-run evolution of rent sharing among UK-domiciled companies.
 - We construct a comprehensive and consistent panel of firms since 1983, spanning the entire economy.
 - Complemented with the analysis of the UK manufacturing firms, and the EU and US industries.
 - Investigating the role of market power (superstar firms).

- We show that UK-domiciled companies share their profits (elasticity .012).
- Decline in rent sharing, the elasticity after 2000 is four-time smaller than before.
- Similar findings for other datasets and countries.
- A positive association between market power and rent sharing, but weaker after 2000.

- One of the first comprehensive studies to estimate the long-run evolution of rent sharing.
 - Bell and Van Reenen (2011) document falling rent sharing for the US manufacturing industries.
 - Benmelech, Bergman and Kim (2018) present similar findings for the US manufacturing companies.
 - Our study covers the entire economy and looks at global and domestic operations.
- ► A decline in rent sharing:
 - \rightarrow growing capital share.
 - \rightarrow falling firm-wage premia.
- The role of market power. Competition policies should also be analysed from the labour market perspective.

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Relationship Between Wages and Rents

- A correlation between wages and economic rents is not a feature of a standard perfect competition model.
- A monopsonistic model with upward sloping labour supply curve.
 - Positive demand shock → wages must rise in order to increase employment.
 - Short-run relationship.
- ► An incentive pay model with risk-averse workers and firms.
 - Sharing of good and bad times.
 - Long-run relationship.
- A bargaining model with rent sharing. model
 - Workers and firms bargain over wages. Workers appropriate a portion of rents.
 - The correlation captures workers' bargaining power.
 - Long-run relationship.

- A company divides its economic rents between the owner (profits) and workers (wages above the market level).
- Workers and firms engage in a Nash bargain, with standard maximization problem

$$max[\theta \ln[(u(w) - u(\overline{w}))n] + (1 - \theta)\ln(\pi)]$$

► FOC implies:

$$w \cong \overline{w} + (\frac{\theta}{1-\theta})\frac{\pi}{n}$$

Existing Empirical Evidence

- Studies have found elasticity within the range of .01-.11. more
- The validity of instrumental variables estimates in this literature remains a contentious issue.
 - Most studies tend to instrument firm-level rents with industry-level rents or shocks (e.g. Card et al., 2014; Estavao and Tevlin, 2003), but the exclusion restriction is not likely to be satisfied (Manning, 2011).
 - Some studies use patents (Van Reenen, 1996; Kline et al., 2017), but the first stage is weak.
 - In general, instrumenting profits increases the estimated elasticity.
- We use GMM and two-period (and before) lags as instruments (Arellano and Bond, 1991). Also report estimates using a leave-out industry measure.

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- Our universe are the largest 300 (by market cap) firms on the London Stock Exchange between **1983-2016**, domiciled and registered in the UK.
 - Except investment, unit and real estate trusts.
 - Except firms, which were in the top 300 for <=2 years.
 - Consider all available years, even when outside the top 300.
- 832 companies, 11478 observations. 95% of the market cap, >7mln employees.
- ► Data: more
 - Manually collected from annual reports (Mergent Archives, Company House).
 - Worldscope, Compustat, Orbis, Fame, Cambridge DTI, Exstat.
- ► We capture *global* operation.

Mean Employment



Real Revenue, Compensation and Profit per Employee



The Rankings of Companies

1983 Market Cap (in mln) British Petroleum 2 General Electric Company Imperial Chemical Industries 4 Marks & Spencer Group British American Tobacco

3

5

Employment

7421

4888

3880

2830

2631

32381

8256

7904

7888

6605

1	British American Tobacco	187173
2	General Electric Company	170865
3	Grand Metropolitan	136297
4	British Petroleum	131600
5	Unilever	127000

Revenue (in mln)

1	British Petroleum
2	Imperial Chemical Industries
3	British American Tobacco
4	Barclays
5	Natnl Westminster Bank

2000 Market Cap (in mln) Vodafone Group 158124 British Petroleum 121844 GlaxoSmithKline 118910 HSBC Holdings 91284 AstraZeneca 59619

Employment

Unilever	295000
Anglo American	249000
Sainsbury	185200
HSBC Holdings	161624
Tesco	152210

Revenue (in mln)	
British Petroleum	97900
Aviva	40244
HSBC Holdings	33182
Unilever	28977
Prudential	28078

2016

Market Cap (in mln)	
HSBC Holdings	130498
British Petroleum	99236
British American Tobacco	86162
GlaxoSmithKline	76695
AstraZeneca	56137

Employment

G4S	592897
Compass Group	527180
Tesco	464520
HSBC Holdings	235175
Sainsbury	181900

Revenue (in mln)

British Petroleum	136100
Legal & General Group	77969
Prudential	71842
HSBC Holdings	60495
Tesco	55917

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$$w_{ijt} = \alpha w_{ij-1} + \sum_{l=0}^{L} \beta_l \pi_{ijt-l} + \sum_{l=0}^{L} \gamma_l U_{t-l} + \sum_{l=0}^{L} \delta_l \overline{w}_{jt-l} + \mu_i + f(time) + \epsilon_{ijt}$$

- w_{ijt} log of compensation per employee for company i, industry j at time t.
- π_{ijt} profit before tax per employee.
- U_t log of nationwide unemployment (ONS).
- \overline{w}_{jt} log of industry average wages (KLEMS).
- ► Endogeneity we take first △ and use lagged levels as instruments (Arellano-Bond).
- We trim the 1/99th percentiles of profits per employee (Card et al. 2014).

The UK-domiciled Companies, 1983-2016

			Depend	ent Variable	: Log w _{ijt}			
	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)
Log <i>w</i> _{ijt-1}	0.477***	0.488***	0.43***	-0.177***	0.478***	0.494***	0.445***	-0.187***
	(0.034)	(0.034)	(0.052)	(0.028)	(0.035)	(0.036)	(0.054)	(0.028)
π/n _{ijt}	0.006***	0.008***	0.01***	0.008***	0.006***	0.008***	0.009***	0.008***
	(0.002)	(0.002)	(0.002)	(0.002)	(0.001)	(0.002)	(0.002)	(0.002)
π/n_{ijt-1}	-	-0.002**	-0.003	0	-	-0.002*	-0.003	0.001
		(0.001)	(0.002)	(0.003)		(0.001)	(0.002)	(0.003)
π/n_{ijt-2}	-	-	0.002	0	-	-	0.002	0.001
			(0.002)	(0.001)			(0.002)	(0.001)
π/n_{ijt-3}	-	-	-0.001	-0.002**	-	-	-0.001	-0.002**
			(0.001)	(0.001)			(0.001)	(0.001)
LR Coefficient	0.011	0.010	0.013	0.006	0.011	0.011	0.013	0.007
	(0.003)	(0.003)	(0.003)	(0.004)	(0.003)	(0.003)	(0.003)	(0.004)
Lester Range	0.158	0.144	0.183	0.093	0.160	0.155	0.182	0.108
Firm-Years	11478	11380	9751	9751	11478	11380	9751	9751
Firms	832	829	731	731	832	829	731	731
Time	Quad	Quad	Quad	Quad	Year FE	Year FE	Year FE	Year FE
Instruments	Lag(2/.)	Lag(2/.)	Lag(2/.)	No	Lag(2/.)	Lag(2/.)	Lag(2/.)	No

Standard errors (in parentheses) clustered at firm level. *** p<0.001, ** p<0.01, * p<0.05

				Depen	dent Variabl	e: Log w _{ijt}				
	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)	(9)	(10)
	1983-2000	2001-2016	1983-1991	1991-2000	2000-2009	2009-2016	1983-1991	1991-2000	2000-2009	2009-2016
Log w _{ijt-1}	0.376***	0.428***	0.620***	0.438***	0.512***	0.253***	0.351*	0.359***	0.597***	0.265***
	(0.086)	(0.062)	(0.161)	(0.077)	(0.057)	(0.083)	(0.183)	(0.129)	(0.085)	(0.098)
π/n _{ijt}	0.017***	0.01***	0.002	0.017***	0.010***	0.004	0.013	0.033***	0.008*	0.005
	(0.004)	(0.003)	(0.006)	(0.003)	(0.003)	(0.003)	(0.021)	(0.009)	(0.005)	(0.006)
π/n_{ijt-1}	0	-0.003	0.014	-0.003	-0.005	0.002	0.014	0.006	-0.006	0.007**
	(0.004)	(0.003)	(0.010)	(0.003)	(0.004)	(0.002)	(0.025)	(0.011)	(0.008)	(0.003)
π/n_{ijt-2}	0.004	0.002	0.003	0.006*	0.002	-0.001	0.014	-0.001	-0.003	-0.008*
	(0.003)	(0.002)	(0.008)	(0.003)	(0.002)	(0.001)	(0.025)	(0.009)	(0.005)	(0.005)
π/n_{ijt-3}	0.006*	-0.002*	-	-	-	-	-	-	-	-
	(0.003)	(0.001)								
LR Coefficient	0.043	0.012	0.050	0.035	0.016	0.007	0.065	0.060	-0.003	0.006
	(0.013)	(0.004)	(0.042)	(0.009)	(0.007)	(0.004)	(0.053)	(0.021)	(0.028)	(0.012)
Lester Range	0.445	0.200	0.486	0.373	0.277	0.104	0.627	0.642	-0.056	0.095
Firm-Years	4719	5032	1,901	3,748	3,437	2,474	1,897	3,748	3,437	2,474
Firms	547	503	404	539	494	379	404	539	494	379
Time	Year FE	Year FE	Year FE	Year FE	Year FE	Year FE				
Instruments	Lag(2/.)	Lag(2/.)	Lag(2/.)	Lag(2/.)	Lag(2/.)	Lag(2/.)	Ind. Profits	Ind. Profits	Ind. Profits	Ind. Profits

Standard errors (in parentheses) clustered at firm level. *** p<0.001, ** p<0.01, * p<0.05

- ► Positive rent sharing, elasticity .012.
- Strong decline since 1980s (.04) until today (.01).
- Robust to the exclusion of small companies, and oil and financial sectors.
- Results not affected by the use of industry-level instruments.
- Similar results for the UK Manufacturing companies with domestic operation (ARD/ABS). more

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 459 US manufacturing industries 1963-2011 from NBER-CES Manufacturing Industry Database.

$$w_{jt} = \alpha w_{jt-1} + \sum_{l=0}^{L} \beta_l \pi_{jt-l} + \sum_{l=0}^{L} \gamma_l U_{t-l} + \sum_{l=0}^{L} \delta_l \overline{w}_{jt-l} + \mu_j + f(time) + \epsilon_{jt}$$

- U_t log of nationwide unemployment (BLS).
- \overline{w}_{jt} log of **2-digit** industry average wages (CPS).

The US Manufacturing Industries, Sub-Periods

	Dependent Variable: Log w _{ijt}					
	(1)	(2)	(3)	(4)	(5)	(6)
	1963-2011	1963-1974	1974-1983	1983-1991	1991-2000	2000-2011
$Log w_{ijt-1}$	0.729***	0.362***	0.606***	0.382***	0.506***	0.508***
	(0.016)	(0.044)	(0.029)	(0.049)	(0.028)	(0.031)
π/n _{ijt}	0.005**	0.037***	0.010***	0.012***	0.008***	0.005***
	(0.002)	(0.009)	(0.004)	(0.003)	(0.003)	(0.002)
π/n_{ijt-1}	0	-0.001	-0.004	-0.010***	-0.005	0.001
	(0.002)	(0.009)	(0.004)	(0.004)	(0.004)	(0.002)
π/n_{ijt-2}	-0.003	-0.010**	-0.001	0.004	-0.002	-0.005*
	(0.001)	(0.005)	(0.003)	(0.002)	(0.003)	(0.003)
π/n_{ijt-3}	0.001	0.009**	-0.001	0.003	0.001	0.001
	(0.001)	(0.004)	(0.003)	(0.003)	(0.003)	(0.002)
LR Coefficient	0.014	0.054	0.013	0.014	0.004	0.005
	(0.005)	(0.019)	(0.012)	(0.004)	(0.008)	(0.004)
Lester Range	0.082	0.174	0.045	0.064	0.019	0.032
Industry-Years	21004	4590	4590	4130	4550	4972
Industries	459	459	459	459	458	452
Time	Year FE	Year FE	Year FE	Year FE	Year FE	Year FE
Instruments	Lag(2/.)	Lag(2/.)	Lag(2/.)	Lag(2/.)	Lag(2/.)	Lag(2/.)

Standard errors (in parentheses) clustered at industry level. *** p<0.001, ** p<0.05

- Strong decline for the US manufacturing since the 1960s (.05) until today (0).
- Similar decline for the EU industries since the 1990s (.002) until today (0).

Lester Range Estimates



Bukowski, Machin & Soskice (2019): Rent Sharing by Country



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- Link between market power and labour share (Benmelech et al., 2018; Autor et al., 2017; Adrjan 2018). Do companies with high market power share more or less of their profits?
- We use the data on the UK-domiciled companies (the Top 300 sample) and measure market power as a firm's revenue and employment share in the sample's industry total.

$$\begin{split} w_{ijt} &= \alpha w_{ij-1} + \sum_{l=0}^{L} \beta_l \pi_{ijt-l} + \sum_{l=0}^{L} \theta_l mshare_{ijt-l} + \sum_{l=0}^{L} \delta_l \overline{w}_{jt-l} + \\ &+ \sum_{l=0}^{L} \gamma_l \pi_{ijt-l} \times mshare_{ijt-l} + \mu_i + \mu_t + \epsilon_{ijt} \end{split}$$

Measures of Market Power



Revenue Share



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► Main results:

- The evidence of rent sharing...
- ...but its magnitude has fallen.
- A positive association between market power and rent-sharing, but weaker after 2000.
- Potential implications:
 - Less inclusive growth.
 - Weaker position of workers (see also robocalypse).
 - More competitive labour market.

Existing Empirical Evidence

- US Industry-level estimates:
 - Elasticity of wages with respect to profits between **.01** and **.06** (Katz and Summers, 1989; Blanchflower et al., 1996; Estavao and Tevlin, 2003).
- UK Firm-level estimates:
 - .07-.09 (Nickell and Wadhwani, 1990; Nickell et al., 1994), .11 (Van Reenen, 1996), .02 -.03 (Hildreth and Oswald, 1997; Hildreth, 1998).
- Employee-employer matched data:
 - Portugal: .03 -.09 (Cardoso and Portela, 2009; Martins, 2009; Card et al., 2016). Italy: .06 .08 (Guiso et al., 2005; Card et al., 2014), .02 -.03 (Hildreth and Oswald, 1997; Hildreth, 1998). Similar elasticities reported for France (Margolis and Salvanes, 2001; Fakhfakh and FitzRoy, 2004), Germany (Guertzgen, 2009) and Sweden (Arai, 2003; Arai and Hayman, 2009; Carlsson, Messina and Skans 2014).

The Number of Stocks Listed on the LSE



Decomposition of the Top 300 Sample



The UK-domiciled Companies, IV



The UK Manufacturing Companies

- One should interpret the above results as evidence for UK-domiciled companies, since many firms in our sample have operations extending beyond the border.
- We complement it with a similar analysis of domestic operations from the panel of UK manufacturing companies from ARD/ABS for 1983-2016.

$$w_{irt} = \alpha w_{ir-1} + \sum_{l=0}^{L} \beta_l \pi_{irt-l} + \sum_{l=0}^{L} \gamma_l U_{rt-l} + \sum_{l=0}^{L} \delta_l \overline{w}_{rt-l} + \mu_i + f(time) + \epsilon_{irt}$$

- *i* stands for firm, r for region and t for time.
- U_{rt} regional unemployment from LFS,
- \overline{w}_{rt} regional average wages from NES/ASHE.

The UK Manufacturing Companies, Sub-Periods

	Dependent Variable: Log wirt						
	(1)	(2)	(3)	(4)	(5)	(6)	
	1983-2016	1983-2016	1983-1991	1991-2000	2000-2009	2009-2016	
Log w _{irt-1}	0.372***	0.370***	0.466***	0.365***	0.174***	0.239***	
	(0.027)	(0.037)	(0.04)	(0.034)	(0.062)	(0.042)	
π/n_{irt}	0.0150***	0.0135***	0.058**	0.042***	0.014*	0.016	
	(0.012)	(0.007)	(0.026)	(0.014)	(0.007)	(0.011)	
π/n_{irt-1}	0.0022	0.00251	-0.013	-0.001	0.009	-0.002	
	(0.01)	(0.006)	(0.022)	(0.012)	(0.007)	(0.008)	
π/n_{irt-2}	-0.00942***	0.00982**	-0.014**	-0.021***	-0.008	-0.004	
	(0.004)	(0.004)	(0.006)	(0.006)	(0.005)	(0.006)	
π/n_{irt-3}	0.00177	0.00159	0.006	0.003	0.012**	-0.005	
	(0.004)	(0.004)	(0.005)	(0.006)	(0.006)	(0.005)	
LR Coefficient	0.015	0.012	0.069	0.037	0.033	0.007	
	(0.008)	(0.008)	(0.054)	(0.03)	(0.015)	(0.02)	
Lester Range	0.18	0.15	0.542	0.329	0.406	0.076	
Firm-Years	27250	27250	13,374	9,164	3,700	3,108	
Firms	2797	2797	2,058	1,606	841	619	
Time	Quad	Year FE					
Instruments	Lag(2/.)	Lag(2/.)	Lag(2/.)	Lag(2/.)	Lag(2/.)	Lag(2/.)	
monuments	Lag(2/.)	Lag(2/.)	Lag(2/.)	Lag(2/.)	Lag(2/.)	Lag(2)	

Standard errors (in parentheses) clustered at firm level. *** p<0.001, ** p<0.01, * p<0.05

- EUKLEMS data allow us to look at domestic operation over the entire economy for the numerous EU countries (AT, DE, DK, ES, FI, FR, IT, NT, UK).
- For each country, the panel consists of 25 years of data for 28 industries.

$$\overline{w}_{jct} - \overline{w}_{jct-l} = \beta_l (\pi_{jct} - \pi_{jct-l}) + FE + \epsilon_{jct}$$

- j stands for industry, c for country and t for time.
- ► Two periods: 1991-2005, 2005-2015
- ► *FE* are industry or country fixed effects.

	Dependent variable : Log w_{ijt} - Log w_{ijt-i}						
	(1)	(2)	(3)	(4)			
		1991	-2005				
$(\pi/n)_{ij2005}$ - $(\pi/n)_{ij1991}$	0.0019*** (0.0001)	0.0015*** (0.0001)	0.0017*** (0.0003)	0.0012***			
Lester Range	5%	4%	5%	3%			
		2005	-2015				
$(\pi/n)_{ij2015}$ - $(\pi/n)_{ij2005}$	-0.0003 (0.0003)	-0.0003 (0.0003)	-0.0001 (0.0002)	-0.0001 (0.0002)			
Lester Range	0%	0%	0%	0%			
Observations	255	255	255	255			
Country FE	No	Yes	No	Yes			
Industry FE	No	No	Yes	Yes			

Standard errors (in parentheses) clustered at industry level. *** p<0.001, ** p<0.05