It takes (more than) a moment: Revisiting the link between firm productivity and aggregate exports

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ECB, Frankfurt, June 25-26, 2015

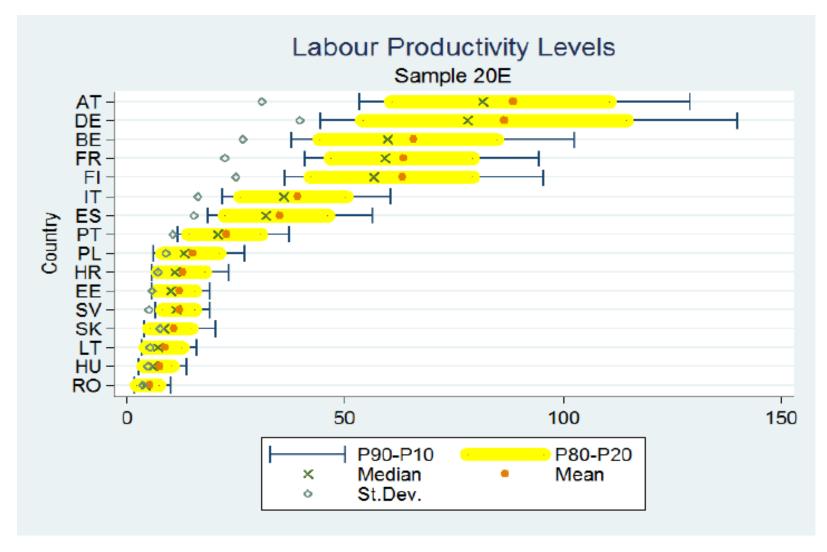
# Motivation

• MICRO TO MACRO:

⇒The aim of this paper is to examine which features of productivity distributions are related to aggregate exports.

- ⇒ (Ir)resistible prominence of average productivity (first moment) in explaining aggregate export
  - $\Rightarrow$  policy practice (e.g. competitivness measured by average unit labour costs)
  - ⇒ micro-macro trade literature (standard trade model à la Melitz 2003; comparative advantage in Ricardian framework à la Costinot et al. 2012)
- $\Rightarrow$  Contradicting hints
  - ⇒ The distribution of firms' characteristics matters for aggregate outcomes (Gabaix, 2011 and Happy Few Mayer Ottaviano, 2007).
  - ⇒ Recent empirical studies provide evidence of the large heterogeneity of firms' performances (TFP and labour productivity) both within and between countries (CompNet).

### **Motivation**



Source: Our calculation from CompNet.

### **Presentation outline**

- What we do and achieve
- Literature
- Theoretical Framework
- Empirical Analysis
- Policy implications

### This paper What does it do?

Which moments of productivity distributions (mean; dispersion; asimmetry) are related to aggregate exports (exporters' multilateral resistance)? Does standard trade theory stand to the data?

- a. Define theoretically the null hypothesis:
  - 1. Derive a gravity equation of the standard trade model à la Melitz (2003), where aggregate bilateral exports from country *o* to any country *d* only depend on the mean productivity (and number of firms) and not on higher moments of the productivity distribution of its producers

#### b. Test empirically the null hypothesis

- 1. THE COMPETITIVENESS INDEX (Multilateral resistance term). Run gravity regressions and derive origin country/industry fixed effects controlling for importer fixed effects and dyadic characteristics.
- 2. RELATE IT TO PRODUCTIVITY DISTRIBUTIONS. Regress these fixed effects on various moments of the country/industry productivity distributions (if higher moments matter too, not only the first moment, the null hypothesis is rejected).

#### This paper What does it achieve?

- 1. The null hypothesis is rejected:
  - a. Aggregate exports are highly correlated to the first moment of productivity distributions, but also to measures of dispersion and asymmetry.
  - b. Asymmetry is especially robust.
- 2. These results hold both for labor productivity and a subsample of indicators TFP.
- 3. The results are robust to different specifications (and standard errors).
- 4. Revisit policy framework

### Literature

- Happy Fews and Granularity.
  - The Happy Few (Clerides Lach and Tybout, 1998, Bernard and Jensen, 1999; Mayer and Ottaviano, 2007): aggregate exports highly concentrated.
  - Granularity (Gabaix, 2011; Di Giovanni & Levchenko, 2013): idiosyncratic shocks to large firms affect aggregate fluctuations.
- New Gravity
  - Structural gravity with heterogeneous firms and Pareto, exporters multilateral resistance terms is only related to exporting country's first moments and number of firms (from Melitz, 2003; Chaney, 2008; Helpman et al.,2008; Head and Mayer, 2014).
- Gains from trade and productivity distributions
  - Pareto to quantify welfare effects of trade by using a parsimonious set of aggregate stats (Arkolakis et al. ,2012; Costinot and Rodriguez-Clare, 2014).
  - Away from Pareto to correctly compute variable bilateral elasticities to trade costs (Bas et al., 2015; Melitz and Redding, 2015).

### **Theoretical Framework**

Aggregate export in generalized trade model with heterogeneous firms

$$X_{od} = \frac{N_o \int_0^{c_{od}} u'(q_{od}(c))q_{od}(c)dG_o(c)}{\sum_{m=1}^M N_m \left[\int_0^{c_{md}} u'(q_{md}(c))q_{md}(c)dG_m(c)\right]} y_d L_d$$

Aggregate export in standard trade model à la Melitz (2003)

- CES demand system
- Iceberg variable and fixed trade costs
- Pareto distribution

$$X_{od}'' = N_o (\overline{c}_{oo})^{-k} (f_{od})^{1 - \frac{k}{\sigma - 1}} (\tau_{od})^{-k} \frac{y_d L_d (\overline{c}_{dd})^{1 - \sigma + k} (f_{dd})^{\frac{k}{\sigma - 1} - 1}}{N_d^s (\overline{c}_d^s)^{1 - \sigma}}$$
(1)  
xporters' capability  
Multilateral  
esistance) Only first  
moments  
matter

### **Theoretical Framework**

• If we explicit factor prices and productivity assuming technology is Cobb Douglas in labour and capital with shares  $\alpha$  and  $(1-\alpha)$  (see Head and Mayer, 2014), exporters' capabilities is:

$$N_{o}(\bar{c}_{oo})^{-k} = N_{o} \left( \alpha^{\alpha} (1-\alpha)^{1-\alpha} \right)^{-k} (w_{o}^{\alpha} r_{o}^{1-\alpha})^{-k} (\varphi_{oo})^{k}$$
Average
productivity

- We assume
  - capital freely mobile within and across sectors (captured by country fixed effect);
  - labour freely mobile only within sectors

### **Empirical Strategy**

We test the implication of the model in two steps.

- 1. We run gravity regressions to estimate origin country fixed effects for a sample of Eurozone countries. Fixed effects ideally measure the "competitiveness" of the sampled countries.
- 2. We check whether the variation in the estimated origin country fixed effects is related to various moments of the distribution of firm productivity.

# **Objective:** Test the null hypothesis of the "standard trade model":

- only the first moment of the productivity distribution matter for competitiveness.

### Data

#### **Productivity measures: CompNet Database**

- The CompNet database is a database of comparable productivity indicators for 17 (variable) EU countries built by members of CompNet using state of the art computation methodologies.
  - All firms 1996-2012
  - > 20 employees 2001-2011 (cross country comparable).
- The data exploits the information contents coming not only from averages, but also from the distribution of firms across several dimensions (e.g. productivity, size, sectors).
  - unweighted average, median, coefficient of variation, 10<sup>th</sup>, 20<sup>th</sup>, 80<sup>th</sup>, and 90<sup>th</sup> percentiles, and skewness
- Main origin source: Central Banks and NSI micro level databases.

#### Data

#### **CompNet Database**

- *Countries*: Austria, Belgium, Croatia, Estonia, Finland, France, Germany, Hungary, Italy, Lithuania, Poland, Portugal, Romania, Slovakia, Slovenia, Spain.
- Sectors: manufacturing sectors at NACE 2-digit rev.2 (with the exclusion of Coke and Petroleum (19) and Tobacco (12))
- *Time period*: 1996-2012/2001-2011 (two overlap the two version of CompNet database).
- Productivity measured as labour productivity (value added per worker) and TFP (Wooldridge with modifications by Galuscak).
- In the second stage, we eliminate observational units that are obtained with less than 10 observations (at least 10 firms by sector, year, and country).
- Unbalanced panel.

#### **Data** Gravity Data

- *Eurostat Comext:* export values in (logs) of millions of euros by destination, country, year & sector.
- *Eurostat*: labor compensation and total employment by sector.
- **Cepii:** distance, common border, common language, former colony.
- **UNCTAD:** nominal bilateral exchange rate.

#### Empirical Analysis First Step

#### Gravity:

- Unbalanced panel of 472,321 observations
- Baseline: includes all bilateral export flows from 20E CompNet countries (o) to destination countries (d) and 22 manufacturing sectors (s) from 2001 to 2012 (t).
- We estimate Eq. (1) as follows

$$Log(Export)_{o,d,s,t} = \alpha_{o,s,t} + \beta_{d,s,t} + \gamma_{o,d} + \varepsilon_{o,d,s,t}$$

- $\alpha_{o,s,t}$  :origin\*year\*sector fixed effects -> Competitiveness index
- $\beta_{d,s,t}$ : destination \* year\*sector fixed effects
- $\delta_{o.d}$ : dyadic terms (distance, common border, etc...)

Fixed effects  $\alpha_{o,s,t}$  measure the competitiveness of the sampled countries as suppliers, netting out importer-specific and country-pair-specific characteristics

**First step results** 

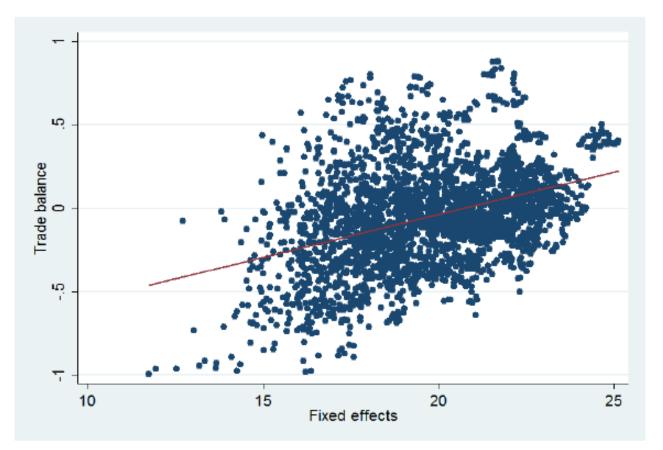
	(1)	(2)	(3)
	All	From 2001	From 2001 (Country in
			sample 20E)
Log(Distance)	-1.219***	-1.235***	-1.214***
	(.0601)	(.064)	(.0761)
Common Border	.5166***	.5775***	.4655***
	(.1476)	(.1527)	(.1546)
Common Language	.7079***	.7123***	.7222***
	(.1005)	(.1006)	(.1315)
Former Colony	.6125***	.5727***	.6496***
	(.1659)	(.1643)	(.1641)
Obs	775764	578965	472321
$\mathbb{R}^2$	.8248	.8185	.8268
Fixed Effects 1	Origin*Sector*Year	Origin*Sector*Year	Origin*Sector*Year
Fixed Effects 2	Destination * Sector * Year	${\rm Destination}^*{\rm Sector}^*{\rm Year}$	Destination * Sector * Year

#### First step: Competitive index, descriptive statistics

Country	Mean	St.Dev	IQR	Obs	Min	Max
Austria	12.642	0.978	0.919	227	9.170	14.575
Belgium	13.132	1.278	1.836	242	9.217	15.327
Croatia	10.070	0.700	0.953	224	7.859	11.457
Estonia	10.285	0.795	0.879	239	7.065	11.775
Finland	12.002	1.247	1.518	250	8.606	14.519
France	13.959	1.298	1.554	252	9.735	15.798
Germany	14.535	1.276	1.698	252	11.088	17.048
Hungary	11.505	1.091	1.495	210	8.160	14.253
Italy	14.161	1.088	0.637	253	10.237	16.612
Lithuania	10.569	0.820	0.804	211	7.661	12.224
Malta	9.671	0.870	1.324	36	8.265	11.042
Poland	12.437	0.974	1.174	176	9.359	14.337
Portugal	12.087	0.691	0.714	147	9.512	13.016
Romania	11.251	0.896	1.119	210	8.519	12.853
Slovakia	11.012	0.934	0.937	229	8.338	13.615
Slovenia	10.976	0.860	1.063	237	8.128	12.601
Spain	13.388	1.023	1.148	252	9.603	14.997
Total	12.159	1.756	2.649	3647	7.065	17.048

- The estimated  $\alpha's$  are highly correlated among all the specifications.
- The most advanced European economies, such as Germany or France, show the highest values, while smaller countries (e.g., Estonia or Romania) reports the lower ones.
- Fixed effects are positively correlated with sectoral trade balance (index of correlation equal to 0.30).

First step: Competitiveness index and trade balance



Source: authors elaborations on Eurostat ComExt data. Each dot is defined at country-sector-year level. The Y-axis reports the trade balance defined as  $\frac{export-import}{export+import}$ . The X-axis reports the fixed effect computed from Eq. 11 (see Table 3, Col.3.). The red line represents the linear interpolation.

#### Second stage: Asimmetry and Dispersion measures

#### Asymmetry

For each country-sector-year triple, we measure the asymmetry of distribution using parametric (*Skewness index – third moment*) and non parametric (*Pearson's second skewness coefficient*) asymmetry indices

 $Pears._{o,s,t} = (\frac{mean_{ost} - median_{dst}}{st.dev._{ost}}),$ 

and the skewness index reported in the CompNet database (third moment).

#### Dispersion

- The ratio of the 80th to the 20th percentile of the productivity distribution (P80/P20).
- The ratio of the 90th to 10th percentile of the productivity distribution (P90/P10).

### Data

#### Second Stage: Productivity distributions from CompNet Database

	Labour Productivity							
Country	LProd(Mean)	LProd(Asim.)	LProd(Skew.)	LProd(P90/P10)	LProd(P80/P20)			
Austria	4.347	0.175	0.907	2.528	1.862			
Belgium	4.004	0.220	1.287	2.725	1.908			
Croatia	2.328	0.249	1.271	4.154	2.522			
Estonia	2.321	0.208	0.976	3.726	2.402			
Finland	4.018	0.218	1.234	2.469	1.753			
France	4.121	0.217	1.207	2.714	1.910			
Germany	4.516	0.200	1.193	3.175	2.108			
Hungary	1.784	0.262	1.572	5.848	2.921			
Italy	3.582	0.205	1.301	2.840	1.950			
Lithuania	2.024	0.269	1.317	5.347	3.113			
Poland	2.436	0.252	1.667	4.897	2.767			
Portugal	2.978	0.202	1.125	3.278	2.163			
Romania	1.497	0.300	1.813	5.814	3.178			
Slovakia	2.263	0.273	1.921	5.555	2.954			
Slovenia	2.392	0.182	1.045	2.967	1.960			
Spain	3.496	0.204	1.127	3.021	2.039			
Total	3.102	0.225	1.307	3.739	2.309			

#### Second Step More than one moment!!!

Following the <u>theoretical model</u>

 $Comp. Ind_{o,s,t} = a_0 + a_1 Asim_{o,s,t-1} + a_2 Dis_{o,s,t-1} + a_3 Mean_{o,s,t-1} + D_o + D_s + D_t + e_{ost}$ 

- The dependent variable is the competitiveness index (*Comp.Ind.*), i.e, the fixed effects of gravity.
- First moment: *Mean* is the average productivity level as computed from the CompNet database.
- Higher moments: *Asim ; Disp* measures of asymmetry and dispersion.
- We lag explanatory variables of one year to minimize concerns of reverse causality.
- We include different combinations of fixed effects (country, sector, and year).

#### Second Step: benchmark labour productivity

	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)	(9)
$Log(LProd(Mean))_{ost-1}$	.0841***	.0804***	.0748***	.0944***	.1056***	.0919***	.0835***	.1021***	.0947***
$Log(firms)_{ost-1}$	(.0232) $.5957^{***}$	(.0231) $.595^{***}$	(.0223) $.5966^{***}$	(.022) .5964***	(.0222) .5772***	(.022) $.5961^{***}$	(.0214) $.597^{***}$	(.0222) $.5776^{***}$	(.0216) $.58^{***}$
0( )000 1	(.0314)	(.031)	(.0303)	(.0314)	(.0319)	(.0312)	(.0304)	(.0317)	(.0308)
$Log(wage)_{ost-1}$	0627*** (0142)	0654*** ( 0144)	$0701^{***}$	0599*** ( 0128)	0598*** ( 014)	0615*** ( 0141)	067*** ( 0120)	0618*** (0142)	0671*** (014)
$\mathrm{LProd}(\mathrm{P90}/\mathrm{P10})_{ost-1}$	(.0142)	(.0144) $.0136^{*}$ (.0079)	(.0142)	(.0138)	(.014)	(.0141) .0072 (.0082)	(.0139)	(.0142) .0093 (.008)	(.014)
$LProd(P80/P20)_{ost-1}$		()	$.1054^{***}$ (.0251)			(.0002)	$.0886^{***}$ (.0251)	(.000)	$.0983^{***}$ (.0249)
$LProd(Pears.)_{ost-1}$				.4133*** (.1177)		$.3913^{***}$ (.1244)	.2887** (.1192)		
$\operatorname{LProd}(\operatorname{Skew.})_{ost-1}$				()	$.0761^{***}$ (.0191)	()	()	$.0728^{***}$ (.0195)	$.0681^{***}$ (.0185)
Cons.	9.784***	9.77***	9.628***	9.634***	9.661***	9.634***	9.548***	9.657***	9.529***
	(.2088)	(.2092)	(.2087)	(.202)	(.2052)	(.2023)	(.2048)	(.2055)	(.2078)
Obs.	2789	2789	2789	2789	2789	2789	2789	2789	2789
R2	.9293	.9294	.93	.9309	.9298	.9298	.9298	.9302	.9309
Country fixed Effects	yes	yes	yes	yes	yes	yes	yes	yes	yes
Sector fixed effects	yes	yes	yes	yes	yes	yes	yes	yes	yes
Year fixed effects	yes	yes	yes	yes	yes	yes	yes	yes	yes
Clustering	Sector-year	Sector-year	Sector-year	Sector-year	Sector-year	Sector-year	Sector-year	Sector-year	Sector-year

# Empirical Analysis Second step: quantifying the impact

Increase of one standard	%Δ Country Competitiveness
deviation in:	
Average Productivity	6.2%
Pears Index	2.5%

Asymmetry has as a positive impact, but relatively smaller than average productivity

#### **Robustness I**

	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)
$Log(LProd(Mean))_{ost-1}$	.0963***	.1068***	.0844***	.0949***	.0932***	.1021***	.0814***	.0903***
	(.0242)	(.0244)	(.0234)	(.0236)	(.0239)	(.0239)	(.0234)	(.0234)
$Log(firms)_{ost-1}$	$.6208^{***}$	$.5973^{***}$	$.6215^{***}$	.601***	$.5974^{***}$	$.5799^{***}$	$.5982^{***}$	$.583^{***}$
	(.0336)	(.034)	(.0326)	(.0329)	(.0326)	(.0332)	(.0316)	(.032)
$Log(wage)_{ost-1}$	0542***	0543***	0611***	0613***	0596***	0597***	0664***	0666***
- ( - )	(.0143)	(.0146)	(.0144)	(.0145)	(.0143)	(.0145)	(.0144)	(.0145)
$LProd(P80/P20)_{ost-1}$			.0902***	.1021***			.0875***	.0965***
			(.0256)	(.0254)			(.0271)	(.027)
$LProd(Pears.)_{ost-1}$	.473***		.345***		.3889***		.2624**	
	(.1193)		(.1206)		(.13)		(.1317)	
$LProd(Skew.)_{ost-1}$	( <i>'</i>	.0858***		.0775***		.0678***	( <i>'</i>	.0599***
( )000 1		(.0199)		(.0193)		(.0211)		(.0205)
Cons.	8.097***	8.216***	8.283***	8.294***	11.3***	11.31***	$11.2^{***}$	11.17***
	(.2748)	(.2776)	(.2166)	(.213)	(.155)	(.1564)	(.1617)	(.1646)
Obs.	2789	2789	2789	2789	2789	2789	2789	2789
R2	.9323	.9324	.9328	.933	.9322	.9322	.9327	.9328
Country X year fixed effects	yes	yes	yes	yes	no	no	no	no
Sector X year fixed effects	no	no	no	no	yes	yes	yes	yes
Sector fixed effects	yes	yes	yes	yes	no	no	no	no
Country fixed effects	no	no	no	no	yes	yes	yes	yes
Clustering	Sector-year	Sector-year	Sector-year	$\mathbf{Sector}$ -year	Sector-year	Sector-year	Sector-year	Sector-year

#### **Robustness II**

	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)
		Alternative	e clustering		WLS			
$Log(LProd(Mean))_{ost-1}$	.0844***	.0949***	.0814***	.0903***	.087***	.0973***	.0893***	.0974***
	(.0222)	(.0208)	(.02)	(.0194)	(.022)	(.0218)	(.0226)	(.0225)
$Log(firms)_{ost-1}$	.6215***	.601***	.5982***	.583***	.631***	.6124***	.6114***	.5979***
- • •	(.0234)	(.023)	(.0236)	(.0233)	(.0343)	(.0346)	(.0334)	(.0337)
$Log(wage)_{ost-1}$	0611***	0613***	0664***	0666***	055***	0553***	0601***	0603***
_ 、 _ ,	(.0147)	(.0144)	(.0133)	(.0128)	(.0138)	(.0139)	(.0137)	(.0138)
$LProd(P80/P20)_{ost-1}$	.0902***	.1021***	.0875***	.0965***	.1027***	.1162***	.0999***	.1106***
	(.0213)	(.0196)	(.0206)	(.0194)	(.0277)	(.0271)	(.0291)	(.0286)
$LProd(Pears)_{ost-1}$	.345***		.2624**		.364***		.283**	
	(.116)		(.1204)		(.1257)		(.1376)	
$LProd(Skew)_{ost-1}$		.0775***		.0599***		.076***		.0577***
		(.0192)		(.0184)		(.0197)		(.0207)
Cons.	8.283***	8.294***	$11.2^{***}$	11.17***	7.117***	7.102***	11.03***	10.99***
	(.1772)	(.1704)	(.3227)	(.3171)	(.2511)	(.2531)	(.1575)	(.1598)
Obs.	2789	2789	2789	2789	2770	2770	2770	2770
$\mathbb{R}^2$	.9328	.933	.9327	.9328	.9359	<mark>.936</mark> 1	.9358	.9359
Country X year fixed effects	yes	yes	no	no	yes	yes	no	no
Sector X year fixed effects	no	no	yes	yes	no	no	yes	yes
Sector fixed effects	yes	yes	no	no	yes	yes	no	no
Country fixed effects	no	no	yes	yes	no	no	yes	yes
Clustering	Country-year	Country-year	Country-year	Country-year	Sector-year	Sector-year	Sector-year	Sector-year

**Robustness III** 

Excluded Country	$Log(LProd(Mean))_{ost-1}$	$Log(firms)_{ost-1}$	$Log(wage)_{ost-1}$	$LProd(Pears.)_{ost-1}$	Obs.	R2
AUT	.0835***	.6282***	0514***	.5687***	2629	.9317
$\operatorname{BEL}$	.1054***	$.5857^{***}$	0665***	.3981***	2595	.9335
CRO	.0931***	.5815***	0658***	.392***	2709	.9282
$\mathbf{EST}$	.0896***	.5842***	0512***	.4943***	2623	.9266
FIN	.1187***	$.5756^{***}$	0562***	.3102**	2589	.9377
$\operatorname{FRA}$	.1059***	.6004***	0641***	.3245***	2558	.9323
GER	.0842***	.6082***	0594***	.4337***	2558	.9171
HUN	.0843***	.5879***	0524***	.3009***	2601	.9308
ITA	.0854***	.572***	0751***	.3711***	2642	.9261
LIT	.0897***	.5768 * * *	0817***	.4035***	2639	.9274
POL	.0951***	.5892***	0598***	.4335***	2697	.9305
PRT	.1206***	.6084***	0481***	.4678***	2663	.9317
ROM	.1071***	.6306***	0542***	.4242***	2602	.9351
SVK	.0761***	.6044***	0661***	.4027***	2586	.9299
SLO	.1051***	.6164***	0497***	.4624***	2586	.9291
SPA	.0805***	.5957***	0608***	.3866***	2558	.9288

Second Step: robustness TFP

# The available data allow us to compute only two cross country comparable statistics on TFP: mean and asymmetry

	(1)	(2)	(3)	(4)	(5)	(6)
$Log(TFP(Mean))_{ost-1}$	.0435**	.0527***	.0427**	.0519***	.0445**	.0535***
	(.0184)	(.0185)	(.0191)	(.0192)	(.0195)	(.0195)
$Log(firms)_{ost-1}$	.6162***	.6167***	.6202***	.6208***	.619***	.6199***
	(.0343)	(.0339)	(.0356)	(.0352)	(.0357)	(.0354)
$Log(wage)_{ost-1}$	0505***	0483***	0515***	0491***	0483***	0461***
	(.015)	(.015)	(.0159)	(.0157)	(.0158)	(.0156)
$TFP(Pears.)_{ost-1}$		.4085***		.3981***		.4124***
		(.111)		(.1178)		(.1253)
Cons.	$10.02^{***}$	9.87***	$10.32^{***}$	10.17***	$11.66^{***}$	$11.52^{***}$
	(.218)	(.2165)	(.1971)	(.2233)	(.1694)	(.1718)
Obs.	2464	2462	2464	2462	2464	2462
R2	.9327	.9332	.9341	.9346	.9351	.9355
Country fixed effects	yes	yes	no	no	yes	yes
Sector fixed effects	yes	yes	yes	yes	no	no
Country fixed effects	yes	yes	no	no	no	no
Country X year fixed effects	no	no	yes	yes	no	no
Sector X year fixed effects	no	no	no	no	yes	yes
Clustering	sector - year					

## **Conclusions**

- We reject the null hypothesis of the standard trade model that it is average productivity that matters for aggregate exports
- The dispersion and the asymmetry of productivity distributions must be taken into account, along with mean productivity, when explaining aggregate export performance.
- Transition towards higher productivity percentiles should be a central policy objective, besides from aiming at the average productivity of the productive sector.